



A clinical risk model for assessing the survival of patients with stage IA–IIA non-small cell lung cancer after surgery

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Background: The survival of patients with stage IA–IIA non-small cell lung cancer (NSCLC) after surgery is heterogeneous. This study aimed to construct a prognostic risk model to predict the overall survival (OS) of these patients.

Methods: Data from patients (n=9,914) from the Surveillance Epidemiology and End Results (SEER) database were analyzed. The cases were randomly divided into the training and the validation groups. Patients from the Shanghai Pulmonary Hospital (n=270) were also included as an external cohort. Independent significant factors affecting survival in the training cohort were used to construct a nomogram. The precision was evaluated using the concordance index (C-index) and calibration plots. The X-tile software was used to confirm the optimal cut-off value to classify the patients.

Results: Sex, age at diagnosis, tumor size, visceral pleura invasion (VPI), tumor grade, and the number of examined lymph nodes were deemed independent prognostic factors and were selected to establish the nomogram. The C-indices of the nomogram for predicting OS were 0.671 [95% confidence interval (CI): 0.653–0.689] in the training group, and 0.668 (95% CI: 0.650–0.687) and 0.707 (95% CI: 0.651–0.763) in the validation and the testing groups, respectively. The cut-off value of risk points was 106.0, which stratified the patients into high-risk and low-risk groups. The high-risk patients had shorter 5-year OS than low-risk patients (P<0.001).

Conclusions: The established nomogram could evaluate the survival in patients with stage IA–IIA NSCLC after surgery and may provide prognostic information for clinicians to make decisions in the management of adjuvant therapy.

Keywords: Non-small cell lung cancer (NSCLC); nomogram; prognosis; survival; treatment

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Introduction

Lung cancer is one of the most common carcinomas globally, accounting for 11.4% of all cancer diagnoses, and it is the leading cause of cancer-related death worldwide

(approximately 18.0%) (1,2). According to pathological types, lung cancer can be divided into small cell lung cancer (SCLC) and non-small cell lung cancer (NSCLC), of which NSCLC accounts for about 80–85% (2,3). For patients with early-stage NSCLC, the main treatment strategy is radical

resection of the primary tumor, which has a 5-year overall survival (OS) rate of 73–90% (3-5). However, there is still a risk of tumor recurrence and metastasis after surgery in a portion of these patients, leading to a decrease in survival after surgical removal of the primary tumor (6,7). This risk is the main reason some patients with early-stage carcinoma require more frequent follow-up management and adjuvant therapy after surgery, which can improve the survival of patients to a certain extent (8,9). Administration of adjuvant therapy is strongly recommended for all patients with advanced NSCLC to improve their prognosis (10), but for patients with early-stage lung cancer, rigorous selection is required. The update of the eighth edition of the American Joint Committee on Cancer (AJCC) Staging Manual has proposed some high-risk factors related to poor survival in early-stage patients (11-13). Many studies have also confirmed that tumor sizes larger than 4 cm and visceral pleura invasion (VPI) could aggravate the prognosis of patients with early-stage lung cancer (14-17).

In the past, clinicians usually evaluated the prognosis of patients through the tumor-node-metastasis (TNM) staging system and used it to guide some options of clinical treatment. Nonetheless, even with NSCLC of the same stage, individuals have significant heterogeneity in survival, so using the TNM staging system alone may not accurately predict survival (18,19). Currently, many tumor-related prognostic models have been established, mainly in the form of nomograms (20-22). A nomogram contains the important prognosis factors and can provide individual risk scores for each person; therefore, it is considered a reliable way to predict the survival of patients (20,22,23). In addition, several studies have suggested that nomograms provide a more accurate survival prediction of various malignancies than traditional TNM staging systems (2,21,24,25). We obtained patient data from the Surveillance Epidemiology and End Results (SEER) database. We aimed to assess the risk factors that could affect the survival of postoperative patients with early-stage NSCLC. A prognostic nomogram was then created to calculate the risk scores of patients. According to the risk scores, patients were classified into high-risk and low-risk groups. The nomogram might provide information for clinicians to evaluate the prognosis of early-stage patients and may guide them in decision-making. We present the following article in accordance with the TRIPOD reporting checklist (available at <https://jtd.amegroups.com/article/view/10.21037/jtd-22-890/rc>).

Methods

Study population

The study was conducted in accordance with the Declaration of Helsinki (as revised in 2013). The study was approved by the Ethics Committee of Shanghai Pulmonary Hospital (No. K22-209). The requirement for informed consent was waived since the SEER database is publicly available and anonymous. Our study cohort consisted of 9,914 patients from the SEER database who underwent surgery between January 2004 and December 2015 and 270 patients from Shanghai Pulmonary Hospital who underwent surgery between 2010 and 2013. Patients who met all of the following selection criteria were enrolled in this retrospective study: (I) histopathologic diagnosis confirmed as NSCLC; (II) no metastasis to the lymph node or other organs; (III) known resection type; (IV) exact differentiation and location of the tumor; (V) the had presence of 1 primary tumor only; (VI) tumor size was between 0 and 5 cm; and (VII) age was within the range of 18 to 80 years. Patients who met any of the following conditions were excluded from this study: (I) a past or current history of another malignancy; (II) invasion of the parietal pleura, vessels, or ribs; and/or (III) they had received neoadjuvant and/or adjuvant chemotherapy. The baseline information of patients we collected covered sex, race, age at diagnosis, surgical approach, tumor grade, histologic type, number of examined lymph nodes, location of the tumor, tumor size, and VPI. Patients from the SEER database were classified into a training cohort (SEER-A) and a validation cohort (SEER-B) by R version 4.1.1 (<https://www.r-project.org/>; The R Foundation for Statistical Computing, Vienna, Austria) to make the baseline data comparable between the 2 groups. The patients from Shanghai Pulmonary Hospital were regarded as an external testing group. The SEER-A cohort and SEER-B cohort each consisted of 4,957 patients, and there was no difference in the baseline data. According to the patients' records, we translated their pathological staging into types set out in the eighth edition of AJCC.

Follow-up and outcome

After the surgical resection of the primary tumor, the patients were followed up. In the SEER database, follow-up duration ranged from 2.0 to 83.0 months, with an average of 59.7 months. Among the patients from Shanghai

Pulmonary Hospital, follow-up duration ranged from 3.0 to 91.0 months, with an average of 68.7 months. The OS was calculated from the date of surgery to the date of death or the last day of follow-up.

Surgery

According to the records in the SEER and Shanghai Pulmonary Hospital databases, the extent of the surgery included sublobar resection, lobectomy, and pneumonectomy. Sublobar resection was implemented when the extent of excision or resection was less than a lobe and included bronchial sleeve resection only, wedge resection, and segmental resection. In the SEER database, the median number of lymph nodes examined was 8.0. In the Shanghai Pulmonary Hospital database, the median number of lymph nodes was 10.0.

Variable declaration

The types of histology were classified as adenocarcinoma (ADC), squamous cell carcinoma (SCC), large cell carcinoma (LCC), acinar cell carcinoma (ACC), adenosquamous carcinoma (ASC), bronchial alveolar carcinoma (BAC), and others. The locations of the primary tumor included the right upper lobe (RUL), right middle lobe (RML), right lower lobe (RLL), left upper lobe (LUL), left lower lobe (LLL), and others. The grades ranged from I to IV and represented well differentiated, moderately differentiated, poorly differentiated, and undifferentiated, respectively.

Statistical analysis

The proportions of categorical outcomes were assessed with the Pearson chi-squared test and Fisher's exact test. OS of patients was analyzed using the Kaplan-Meier method, and the difference in survival between the 2 groups was compared using a log-rank test. Univariable and multivariable Cox proportional hazards models were adopted to identify the independent prognostic predictors. Predictors ($P < 0.05$) in the univariable analysis and known factors affecting prognosis were brought into a multivariable analysis. Results of the univariable and multivariable analyses were presented as hazard ratio (HR) and 95% confidence interval (CI), respectively, and the 2-sided statistical significance level was considered $P < 0.05$. Statistical analyses were conducted using SPSS 23.0 (IBM

Corp., Armonk, NY, USA). The nomogram was constructed with R version 4.1.1 based on the risk factors concluded from the multivariable analysis. The concordance index (C-index) of the model was measured by comparing the predicted survival with the observed survival probability: the larger the C-index, the more accurate the prognostic stratification. The cut-off values for total risk points were assessed using the X-tile software (Copyright: Camp/Rimm; Yale University, New Haven, CT, USA), and patients were divided into high-risk and low-risk groups based on the cut-off value of their risk score.

Results

Patients and characteristics

After selection, 9,914 patients during 2004–2015 were finally included in this study from the SEER database. Patients selected were randomly divided into the SEER-A and SEER-B groups. Among the 9,914 patients, 4,353 (43.9%) were men and 5,561 (56.1%) were women. There were 1,016 (10.2%) patients with VPI, and the others did not have VPI. Patients' ages ranged from 19 to 80 years (median, 65 years; *Table 1*). The primary tumors were located in RUL (35.8%), RML (5.5%), RLL (18.1%), LUL (25.7%), and LLL (14.0%), and the remaining portion of tumors were in other locations ($n=92$, 0.9%) including the main bronchi and multiple positions. The main histologic types were ADC ($n=6,240$, 63.0%) and SCC ($n=2,334$, 23.5%). A total of 4,827 (48.7%) patients had less than or equal to 7.0 lymph nodes removed, and 5,087 (51.3%) patients had more than 7.0 lymph nodes removed during the surgery. Regarding the grade of tumor differentiation, the tumors in 2,356 (23.8%) patients were well differentiated, 4,864 (49.1%) were moderately differentiated, 2,600 (26.2%) were poorly differentiated, and 94 (0.9%) were undifferentiated. In the training cohort, the 1-, 3-, and 5-year OS rates were 95.0%, 84.5%, and 74.3%, respectively. In addition, in the validation cohort, the 1-, 3-, and 5-year OS rates were 94.9%, 84.2%, and 74.9%, respectively. Clinical characteristics of patients in Shanghai Pulmonary Hospital are listed in *Table 2*, and their 1-, 3-, and 5-year OS rates were 94.0%, 82.8%, and 69.7%, respectively.

Univariable and multivariable analyses

To determine the independent prognostic factors for the survival of patients after surgery, we analyzed the

Table 1 The characteristics of patients in the SEER database

Variables	All patients (n=9,914)	Training cohort (SEER-A; n=4,957)	Validation cohort (SEER-B; n=4,957)	P value
Sex, n (%)				
Male	4,353 (43.9)	2,180 (44.0)	2,173 (43.8)	0.887
Female	5,561 (56.1)	2,777 (56.0)	2,784 (56.2)	
Age at diagnosis (years), n (%)				
≤65	4,202 (42.4)	2,133 (43.0)	2,069 (41.7)	0.193
>65	5,712 (57.6)	2,824 (57.0)	2,888 (58.3)	
Location of tumor, n (%)				
RUL	3,548 (35.8)	1,803 (36.4)	1,745 (35.2)	0.815
RML	545 (5.5)	270 (5.4)	275 (5.5)	
RLL	1,795 (18.1)	874 (17.7)	921 (18.6)	
LUL	2,544 (25.7)	1,271 (25.6)	1,273 (25.7)	
LLL	1,390 (14.0)	693 (14.0)	697 (14.1)	
Others	92 (0.9)	46 (0.9)	46 (0.9)	
Tumor size (cm), n (%)				
≤1	821 (8.3)	415 (8.4)	406 (8.2)	0.883
1–2	4,173 (42.1)	2,084 (42.0)	2,089 (42.1)	
2–3	3,071 (31.0)	1,539 (31.1)	1,532 (30.9)	
3–4	1,347 (13.6)	660 (13.3)	687 (13.9)	
4–5	502 (5.0)	259 (5.2)	243 (4.9)	
Histologic type, n (%)				
ADC	6,240 (63.0)	3,089 (62.3)	3,151 (63.6)	0.049
SCC	2,334 (23.5)	1,213 (24.5)	1,121 (22.6)	
ASC	196 (2.0)	110 (2.2)	86 (1.7)	
ACC	568 (5.7)	279 (5.6)	289 (5.8)	
BAC	131 (1.3)	67 (1.4)	64 (1.3)	
LCC	151 (1.5)	70 (1.4)	81 (1.7)	
Others	294 (3.0)	129 (2.6)	165 (3.3)	
Tumor grade, n (%)				
I	2,356 (23.8)	1,194 (24.1)	1,162 (23.4)	0.658
II	4,864 (49.1)	2,429 (49.0)	2,435 (49.2)	
III	2,600 (26.2)	1,292 (26.1)	1,308 (26.4)	
IV	94 (0.9)	42 (0.8)	52 (1.0)	
VPI, n (%)				
No	8,898 (89.8)	4,435 (89.5)	4,463 (90.0)	0.354
Yes	1,016 (10.2)	522 (10.5)	494 (10.0)	

Table 1 (continued)

Table 1 (continued)

Variables	All patients (n=9,914)	Training cohort (SEER-A; n=4,957)	Validation cohort (SEER-B; n=4,957)	P value
Surgical approach, n (%)				
Sublobar resection	1,403 (14.2)	698 (14.1)	705 (14.2)	0.586
Lobectomy	8,411 (84.8)	4,214 (85.0)	4,197 (84.7)	
Pneumonectomy	100 (1.0)	45 (0.9)	55 (1.1)	
Examined lymph nodes, n (%)				
≥8	5,087 (51.3)	2,514 (50.7)	2,573 (51.9)	0.236
≤7	4,827 (48.7)	2,443 (49.3)	2,384 (48.1)	
Race, n (%)				
White	8,129 (82.0)	4,072 (82.1)	4,057 (81.8)	0.104
Black	912 (9.2)	429 (8.7)	483 (9.8)	
Others	832 (8.4)	438 (8.8)	394 (7.9)	
Unknown	41 (0.4)	18 (0.4)	23 (0.5)	

P value was calculated by χ^2 test. SEER, Surveillance Epidemiology and End Results; RUL, right upper lobe; RML, right middle lobe; RLL, right lower lobe; LUL, left upper lobe; LLL, left lower lobe; VPI, visceral pleural invasion; ADC, adenocarcinoma; SCC, squamous cell carcinoma; LCC, large cell carcinoma; ACC, acinar cell carcinoma; ASC, adeno-squamous carcinoma; BAC, bronchial alveolar carcinoma.

OS using a Cox regression model. The results of the univariable and multivariable analyses are shown in *Table 3*. The multivariable analysis confirmed the following 6 characteristics as independent prognostic factors: sex (HR 0.706; 95% CI: 0.619–0.806; $P < 0.001$), age at diagnosis (HR 1.676; 95% CI: 1.453–1.933; $P < 0.001$), tumor size (2–3 cm: HR 1.394; 95% CI: 1.037–1.875; $P = 0.028$. 3–4 cm: HR 1.528; 95% CI: 1.110–2.104; $P = 0.009$. 4–5 cm: HR 1.772; 95% CI: 1.238–2.535; $P = 0.002$), tumor grade (II: HR 2.333; 95% CI: 1.855–2.934; $P < 0.001$. III: HR 2.563; 95% CI: 2.004–3.280; $P < 0.001$. IV: HR 2.769; 95% CI: 1.440–5.321; $P = 0.002$), the number of examined lymph nodes (HR 0.867; 95% CI: 0.761–0.988; $P = 0.032$), and VPI (HR 1.349; 95% CI: 1.120–1.624; $P = 0.002$).

Construction and validation of the prognostic risk model

First, we built a nomogram for OS that included the independent prognostic factors mentioned above (*Figure 1*). The nomogram showed that tumor grade made the largest contribution to prognosis. We could easily draw a straight line down to determine the estimated probability of survival at each time point through the nomogram. The calibration curves showed a good agreement between the nomogram-predicted probability and the actual observation for 3- and

5-year OS (*Figure 2A,2B*). The C-index of the nomogram for predicting OS was 0.671 (95% CI: 0.653–0.689). We counted the scores of all patients and confirmed the optimal cut-off value of the scores by X-tile software. Cases were classified into high-risk and low-risk groups according to the optimal cut-off value (*Figure 3*). Then, the OS of those 2 groups in the validation cohort and Shanghai Pulmonary Hospital cohort were compared using the Kaplan-Meier method. Patients in the low-risk subgroups showed statistically better survival than those in the high-risk subgroups ($P < 0.001$; *Figure 4A,4B*).

Discussion

In this study, based on the results of univariable and multivariable analyses, factors such as age at diagnosis, sex, tumor grade, the number of examined lymph nodes, tumor size, and VPI were found to influence the prognosis in NSCLC patients in stages IA to IIA. We constructed a nomogram to improve the prediction of survival. The calibration plots showed optimal agreement between the prediction and actual observation. The C-indices of the training cohort (0.671), validation cohort (0.668), and testing cohort (0.707) also demonstrated the reliability of the established nomogram. Further, patients were

Table 2 The characteristics of patients in Shanghai Pulmonary Hospital

Variables	All patients (n=270)
Sex, n (%)	
Male	168 (62.2)
Female	102 (37.8)
Age at diagnosis (years), n (%)	
≤65	170 (63.1)
>65	100 (36.9)
Location of tumor, n (%)	
RUL	77 (28.5)
RML	20 (7.4)
RLL	48 (17.8)
LUL	84 (31.1)
LLL	25 (9.3)
Others	16 (5.9)
Tumor size (cm), n (%)	
≤1	33 (12.2)
1–2	92 (34.1)
2–3	82 (30.4)
3–4	42 (15.5)
4–5	21 (7.8)
Histologic type, n (%)	
ADC	189 (70.0)
SCC	81 (30.0)
Tumor grade, n (%)	
I	54 (20.0)
II	172 (63.7)
III	44 (16.3)
IV	0 (0)
VPI, n (%)	
No	228 (84.4)
Yes	42 (15.6)
Surgical approach, n (%)	
Sublobar resection	69 (25.6)
Lobectomy	188 (69.6)
Pneumonectomy	13 (4.8)
Examined lymph nodes, n (%)	
≤9	170 (63.1)
≥10	100 (36.9)

RUL, right upper lobe; RML, right middle lobe; RLL, right lower lobe; LUL, left upper lobe; LLL, left lower lobe; VPI, visceral pleural invasion; ADC, adenocarcinoma; SCC, squamous cell carcinoma.

categorized into high-risk and low-risk subgroups with significant differences in OS by the cut-off value. It is well known that tumor size is an important independent prognostic factor; it has now been introduced into the TNM staging system (26). In this study, tumor size was significantly associated with the survival of patients ($P<0.001$). Moreover, poorly differentiated malignancies are always more invasive than well-differentiated malignancies, and the grade of the tumor is well known to be an important independent prognostic factor (27). In this study, tumor grade was significantly associated with OS of patients with IA–IIA NSCLC ($P<0.001$).

Currently, the total morbidity and mortality of NSCLC remain high in China (28). With the widespread application of low-dose computed tomography (LDCT) (1,29,30), the diagnostic rate of the early-stage disease has increased, but the follow-up management and the determination of whether to use adjuvant therapy remain controversial (31). Thus, it is necessary to analyze the prognosis of early-stage patients and identify high-risk patients, then to provide them with different recommendations for postoperative management and adjuvant treatment. Currently, the primary treatment for patients with IA–IIIB remains complete surgical resection of the tumor; however, postoperative patients still have the possibility of tumor recurrence and distant metastasis, which causes heterogeneity in the survival of individuals (31). Therefore, the decision to administer adjuvant therapy to the patients after surgery is always worth considering.

On the one hand, some studies suggest that adjuvant therapy might be an important strategy to reduce recurrence, prolong survival, and improve the quality of life in patients with early-stage NSCLC after the complete removal of tumors (32). On the other hand, other studies have indicated that the administration of adjuvant therapy to patients with stage IA–IIA NSCLC is not always effective and might harm some patients after surgery (33). Guidelines have already indicated that patients with IIB–IIIB NSCLC need to accept adjuvant therapy routinely after surgery, and survival benefits have been supported (34). For postoperative patients with IB–IIA NSCLC, adjuvant therapy is only recommended among patients with risk factors, such as a tumor size larger than 4 cm, VPI, and poor differentiation (35). However, some risk factors can also greatly influence long-term survival, even in patients with IA NSCLC (36,37). Thus, it is necessary to identify the patients with poor prognosis in stage IA–IIA diseases and provide additional therapy or intensive follow-up. The

Table 3 The univariable and multivariable analyses for mortality

Variables	Univariable analysis			Multivariable analysis		
	HR	95% CI	P value	HR	95% CI	P value
Sex						
Male vs. female	0.623	0.547–0.709	<0.001	0.706	0.619–0.806	<0.001
Age at diagnosis (years)						
≤65 vs. >65	1.037	1.028–1.045	<0.001	1.676	1.453–1.933	<0.001
Location of tumor			0.397			
RUL	1.275	0.603–2.695	0.525			
RML	0.961	0.431–2.141	0.922			
RLL	1.265	0.594–2.695	0.543			
LUL	1.288	0.608–2.731	0.509			
LLL	1.108	0.517–2.375	0.792			
Others	1	Reference				
Tumor size (cm)			<0.001			<0.001
4–5	2.563	1.799–3.651	<0.001	1.772	1.238–2.535	0.002
3–4	1.987	1.448–2.727	<0.001	1.528	1.110–2.104	0.009
2–3	1.715	1.277–2.303	<0.001	1.394	1.037–1.875	0.028
1–2	1.131	0.841–1.522	0.415	1.030	0.765–1.387	0.847
≤1	1	Reference		1	Reference	
Histologic type			<0.001			0.005
ADC	1.256	0.763–2.066	0.371	1.507	0.963–2.360	0.955
SCC	2.299	1.392–3.797	0.001	1.722	1.094–2.710	0.437
ASC	2.097	1.143–3.848	0.017	1.489	0.839–2.643	0.687
ACC	1.170	0.635–2.155	0.614	1.292	0.738–2.259	0.590
BAC	0.506	0.198–1.294	0.155	1.291	0.564–2.956	0.474
LCC	3.690	2.011–6.770	<0.001	2.069	1.201–3.566	0.310
Others	1	Reference		1	Reference	
Tumor grade			<0.001			<0.001
IV	4.503	2.521–8.042	<0.001	2.769	1.440–5.321	0.002
III	3.552	2.825–4.467	<0.001	2.563	2.004–3.280	<0.001
II	2.755	2.208–3.437	<0.001	2.333	1.855–2.934	<0.001
I	1	Reference		1	Reference	
VPI						
No vs. yes	1.595	1.328–1.916	<0.001	1.349	1.120–1.624	0.002

Table 3 (continued)

Table 3 (continued)

Variables	Univariable analysis			Multivariable analysis		
	HR	95% CI	P value	HR	95% CI	P value
Surgical approach			0.029			
Sublobar resection	0.569	0.334–0.972	0.039			
Lobectomy	0.518	0.311–0.864	0.012			
Pneumonectomy	1	Reference				
Examined lymph nodes						
≤7 vs. ≥8	1.117	0.981–1.272	0.096	0.867	0.761–0.988	0.032
Race			0.017			
White	2.781	0.391–19.766	0.307			
Black	2.462	0.342–17.706	0.371			
Others	1.858	0.257–13.421	0.539			
Unknown	1	Reference				

HR, hazard ratio; CI, confidence interval; RUL, right upper lobe; RML, right middle lobe; RLL, right lower lobe; LUL, left upper lobe; LLL, left lower lobe; VPI, visceral pleural invasion; ADC, adenocarcinoma; SCC, squamous cell carcinoma; LCC, large cell carcinoma; ACC, acinar cell carcinoma; ASC, adenosquamous carcinoma; BAC, bronchial alveolar carcinoma.

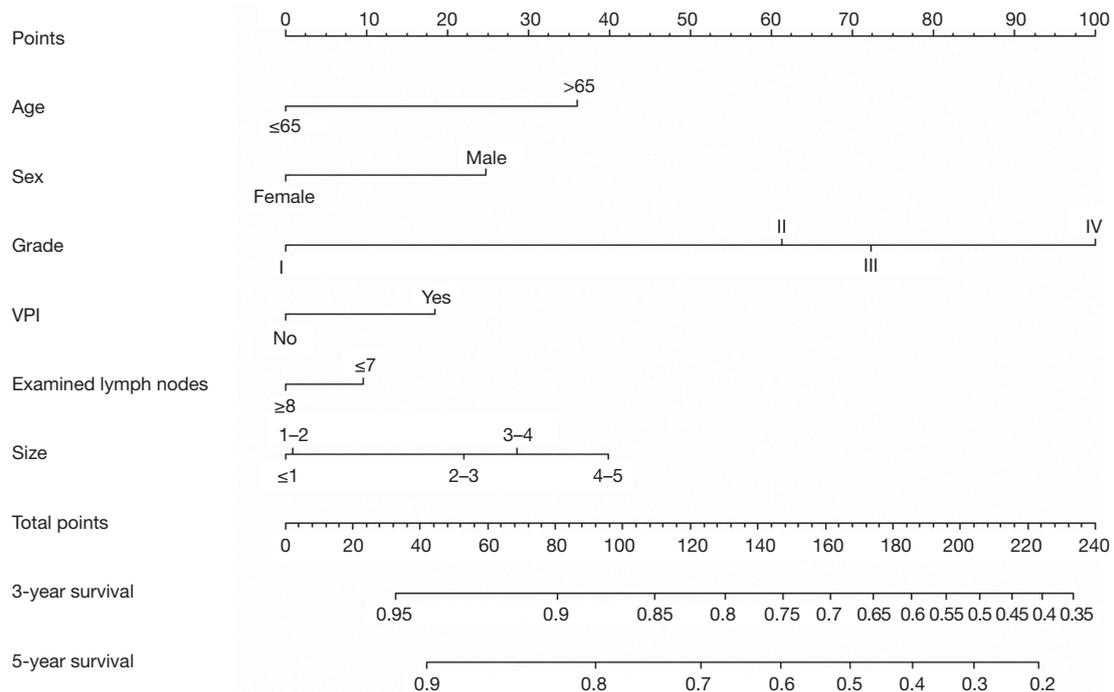


Figure 1 The nomogram for 3- and 5-year OS in patients with stages IA–IIA NSCLC. VPI, visceral pleural invasion; OS, overall survival; NSCLC, non-small cell lung cancer.

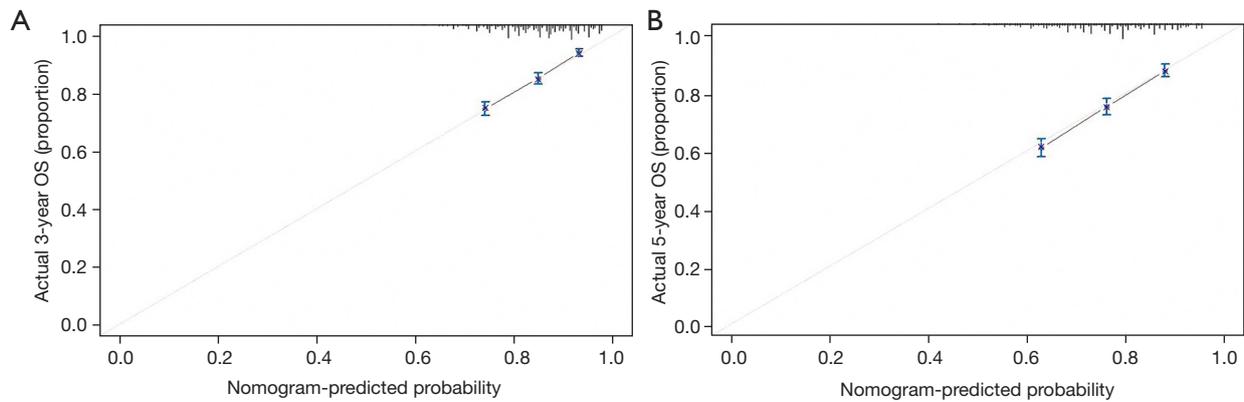


Figure 2 Calibration curves predicting the (A) 3-year, and (B) 5-year OS of patients in the training cohort. OS, overall survival.

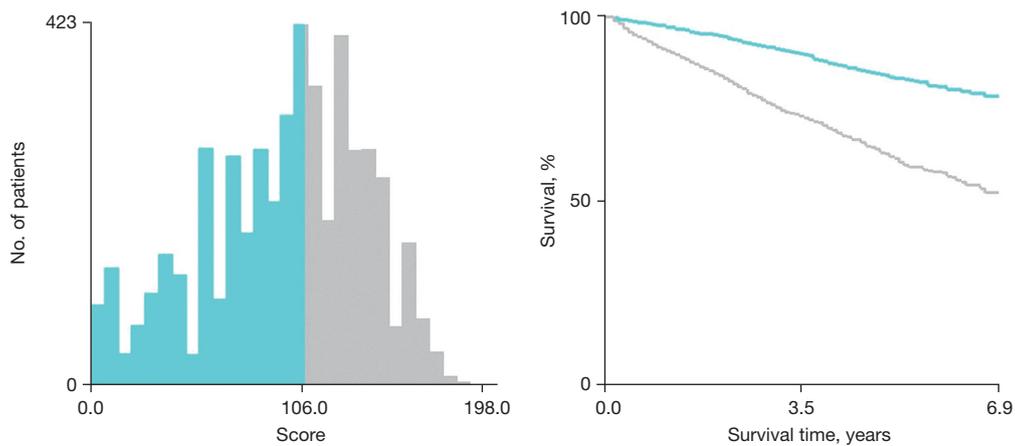


Figure 3 The cut-off value of the risk points counted by X-tile and Kaplan-Meier survival curves for OS in the training cohort. (points ≤ 106.0 were the low-risk group; points > 106.0 were the high-risk group; $P < 0.001$). OS, overall survival.

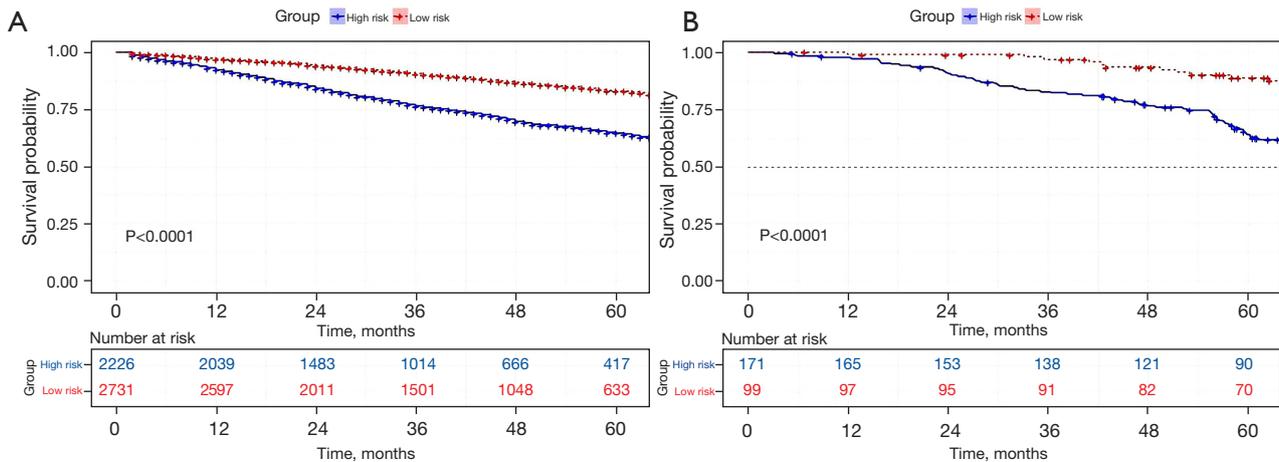


Figure 4 Kaplan-Meier survival curves for OS in the 2 cohorts. Kaplan-Meier survival curves for OS in (A) the validation cohort (SEER-B), and (B) Shanghai Pulmonary Hospital cohort. OS, overall survival; SEER, Surveillance Epidemiology and End Results.

nomogram in this study can successfully identify high-risk patients in a cohort of stage IA–IIA.

Recent years have seen an increasing number of researchers participating in this field, but most previous studies established Cox regression models to analyze the prognosis of patients (38). Those models could not be used in clinical practice due to their low predictive ability and limited extrapolation (39). Several studies have suggested that nomograms provide a more accurate survival prediction of various malignancies than traditional TNM staging systems (24,25). There have been some nomograms created for early-stage NSCLC, yet previous models could not focus on predicting the survival in patients with a tumor size less than or equal to 4 cm, or these earlier models only included a small sample of patients with stage I NSCLC (3,12). Our study incorporated a large sample to build a predictive model for patients with IA–IIA NSCLC. Even though the data from the SEER database may have regional limits, the model was validated by the external testing data from Shanghai Pulmonary Hospital, implying its potential broad applicability.

Despite the precision of the model in our study, there were still some limitations. First, the study was retrospective in nature and failed to incorporate some baseline information of patients from the SEER database, such as the history of smoking, preoperative and postoperative complications, and physical conditions. Second, the prognostic model constructed in this study was based on the data from the SEER database. It is not easy to find a large-scale external validation cohort because of diverse standards in different regions. Third, some pathological factors affecting survival were not included in the SEER database, such as percentages of micropapillary, solid pattern, and the presence of spread through air spaces. Some molecular factors which are potentially important factors affecting the survival of patients with IA–IIA NSCLC were also vacant, such as the mutations of epidermal growth factor receptor and Kirsten rat sarcoma viral oncogene, and the fusion of echinoderm microtubule-associated protein-like 4-anaplastic lymphoma kinase (40–42). Further efforts on prospective data collection and the incorporation of more potential risk factors mentioned above will improve the model.

Conclusions

The nomogram established in this study could predict the prognosis of patients with stage IA–IIA NSCLC after

surgical resection and help clinicians to identify high-risk populations.

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Footnote

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Ethical Statement: The authors are accountable for all aspects of the work in ensuring that questions related to the accuracy or integrity of any part of the work are appropriately investigated and resolved. The study was conducted in accordance with the Declaration of Helsinki (as revised in 2013). The study was approved by the Ethics Committee of Shanghai Pulmonary Hospital (NO. K22-209). The requirement for informed consent was waived since the SEER database is publicly available and anonymous.

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